A

bnormalities of attention have long been thought to be a central feature of schizophrenia. The general function of attention is to provide a competitive advantage when multiple sensory inputs, thoughts, or action plans compete with each other for access to the limited resources of neural representation, awareness, and motor production (1). Basic cognitive neuroscience demonstrates that the brain implements attentional functions by means of a large network of partially independent subsystems (2,3). Thus, one of the major challenges facing the field is to specify which types of attentional mechanisms are impaired in people with schizophrenia (PSZ) and which may be spared (4).

Some types of stimuli or response tendencies have an intrinsic competitive processing advantage and will tend to dominate behavior. For example, high-salience stimuli have a bottom-up processing advantage over low-salience stimuli (5), and high-probability and automated responses have an advantage over low-probability and less automatic responses (6). Attentional control systems are challenged when the less potent stimulus or response tendency is task-appropriate and strong top-down bias signals are necessary to overcome the more potent stimulus or response (7). Thus, impairments in attentional control may be apparent primarily under conditions in which high-potency stimuli or responses must be suppressed.

Indeed, the most persuasive evidence of impaired attentional control in PSZ has been obtained in tasks that emphasize competition between low-potency correct responses and high-potency incorrect responses. For example, deficits have frequently been documented in context versions of the Continuous Performance Test (8), in the Stroop Interference Test (9), and in antisaccade tasks (10,11). In each of these cases, top-down control is needed to inhibit a prepotent response tendency, whether this response has achieved prepotency through task contingencies (context Continuous Performance Test), from preexperiential experience (Stroop Interference Test), or through the reflexivity of eye movements toward sudden-onset stimuli (antisaccade paradigm).

In the basic science literature, most studies of selective attention focus on the selective processing of competing inputs rather than on resolving response competition. However, most of the evidence for impaired attentional control in PSZ has been obtained in tasks that emphasize response selection. There is a paucity of evidence that top-down control processes function abnormally in the selective encoding of visual inputs for further perceptual processing or working memory (WM) storage (for suggestive evidence concerning auditory selective attention, see [12,13]). This may indicate that attentional dysfunction in schizophrenia primarily impacts output-related rather than input-related processes. Alternatively, this lack of evidence may originate from the fact that studies of input selection in PSZ have not typically used tasks in which a low-potency input must be selected in the face of competition from a high-potency input.

Consider, for example, the Posner orienting paradigm (14), in which a cue indicates that attention should be directed to a specific location, and the effectiveness of attentional selection is assessed by comparing performance for targets presented at the cued versus uncued location. With few exceptions (15,16), attentional control processes are minimally challenged in this
paradigm because there is no need to actively ignore irrelevant stimuli. Accordingly, the effectiveness of attentional selection is not found to be impaired in schizophrenia. That is, despite overall slowed responding, the difference between valid and invalid trials is not reduced in PSZ compared with healthy control subjects (HC), indicating that PSZ are able to select spatial locations based on the cue information reviewed by [17].

Similarly, no evidence for impaired attentional selection in PSZ was obtained when cues were used to indicate which stimuli should be encoded into visual WM (18). In four separate experiments, equal numbers of relevant and irrelevant stimuli were presented, and subjects were instructed to remember, for example, the colors of the circles and not the colors of the rectangles. The selection criteria never changed between trials, cues were always simple and salient (shape, color, location), and cued and uncued items were of similar bottom-up salience. Thus, selection of relevant items may have required only a modest boost from top-down bias signals. Accordingly, patients showed robust storage of relevant items and minimal encoding of irrelevant distractors, that is, no evidence of impaired attentional selection.

The present study was designed to test whether PSZ would exhibit specific impairment in attentional selection for WM encoding when control aspects of attentional selection were challenged by irrelevant stimuli holding a competitive salience advantage relative to the relevant stimuli. A deficit, if observed under these conditions, would demonstrate that schizophrenia involves a general deficit in overcoming competition at both encoding and response stages rather than a limited impairment at the stage of response selection.

Methods and Materials

Participants

Twenty-eight patients meeting Diagnostic and Statistical Manual of Mental Disorders (19) criteria for schizophrenia (n = 13 paranoid, 7 undifferentiated, 2 residual) or schizoaffective disorder (n = 6) and 25 matched healthy control subjects participated in this study. Diagnosis was established using a best estimate approach in which information from a Structured Clinical Interview for DSM-IV was combined with a review of patient medical records at a consensus diagnosis meeting chaired by one of the authors (J.M.G.). Demographic information is summarized in Table 1. Groups did not differ in age [t(51) = 2.0, p > .05], parental education [t(48) = .04, p > .4], sex (chi-square p > .9), or ethnicity (chi-square p > .24). However, patients had significantly fewer years of education than control subjects [t(51) = 3.46, p < .01].

The patients were clinically stable outpatients. At the time of testing, patients obtained a total score of 36.1 ± 7.3 (mean ± SD) on the Brief Psychiatric Rating Scale (range 24–53) (20), 36.0 ± 14.3 on the Scale for the Assessment of Negative Symptoms (range 4–73) (21), 20.9 ± 6.3 on the Level of Functioning Scale (range 11–34) (22), and 2.5 ± 2.4 on the Calgary Depression Scale (range 0–9) (23). All patients were receiving antipsychotic medication at time of testing: 4 were treated with first-generation antipsychotics, 22 with second-generation antipsychotics, and 2 with both. Fourteen patients additionally received mood stabilizing medication, five anxiolytic medication and three antiparkinsonian medication. Medication had not changed in the preceding 4 weeks. Control participants were recruited from the community via random digit dialing and word of mouth and had no Axis I or II diagnoses as established by a Structured Clinical Interview for DSM-IV, had no self-reported family history of psychosis, and were not taking any psychotropic medication. Participants provided informed consent for a protocol approved by the University of Maryland School of Medicine Institutional Review Board. Before participants signed the consent form, the investigator reviewed its content with the volunteer and answered any questions. Before volunteers with schizophrenia signed the consent form, the investigator, in the presence of a third-party witness, also formally evaluated basic understanding of study demands, risks, and what to do if experiencing distress or to end participation.

Neuropsychological Testing

Participants completed the Wechsler Abbreviated Scale of Intelligence (24), the Wide Range Achievement Test Reading 4 (25), the Wechsler Test of Adult Reading (26), and the MATRICS Consensus Cognitive Battery (27). Neuropsychological testing was usually performed on a separate day to avoid fatigue. The PSZ scored lower than HC on the Wechsler Abbreviated Scale of Intelligence (p < .001, independent-samples t test) and exhibited significant impairment in all MATRICS Consensus Cognitive Battery domains except visual learning. There were no group differences on the Wide Range Achievement Test Reading 4 (p > .6) or Wechsler Test of Adult Reading (p > .5), suggesting similar premorbid functioning in the two groups (Table 1).

Stimuli

The task was presented in a dimly illuminated room (1.0 foot-candle) on a 17-inch cathode ray tube monitor with a 60-Hz refresh rate. Stimuli were presented against a gray background (luminance approximately 22 candelas/m²). A white fixation cross was presented in the center of the screen throughout each task trial until the trial was ended by a response.

Figure 1 illustrates the sequence of events in each task trial. Each trial started with a 300-msec presentation of six black and white checked (4 × 4) squares, each subtending 98° of visual

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AA, African American; A, Asian; C, Caucasian; O, Other; WASI, Wechsler Abbreviated Scale of Intelligence; WRAT 4, Wide Range Achievement Test Reading 4; WTAR, Wechsler Test of Adult Reading; MATRICS, MATRICS Consensus Cognitive Battery.

*Average over mother’s and father’s years of education.
*Data unavailable for three subjects.
*Data unavailable for one subject.
*p < .01; significant difference between people with schizophrenia and healthy control subjects in independent samples t test.
*p < .001; significant difference between people with schizophrenia and healthy control subjects in independent samples t test.
angle (encoding array). For three of these squares, the black and white checks reversed (in phase across all three squares) at a frequency of 7.5 Hz during the 300-msec stimulus presentation, creating the appearance of flickering. After a 2000-msec delay, all squares reappeared (retrieval array). This time none of them flickered. One square was outlined by a red frame (the probed item). The task consisted of making a forced choice response on one of two buttons to indicate whether the probed item was in exactly the same position as in the encoding array or if it had shifted slightly (by 1.6°) in any direction. On half the trials, the probed item changed position; on the other half, it stayed in the same position. The critical comparison was of memory performance on trials in which the probed stimulus was one of the three items that had been flickering during encoding and trials in which the probed stimulus had not been flickering. The retrieval array stayed on the display until a response was made, followed by a 1-sec intertrial interval during which the screen was blank.

There were three task conditions. The unbiased condition was always run first to avoid any prior attention instructions carrying over and preventing a truly unbiased attentional focus. Here, flickering and nonflickering items were probed with equal probability in a random sequence. This condition consisted of 80 trials, split into two blocks that were separated by a rest period. The flicker-bias and nonflicker-bias conditions were run next, in counterbalanced order. Both of these bias conditions consisted of 200 trials, split into five blocks that were separated by rest periods. In the flicker-bias condition, a previously flickering item was probed on 80% of trials and a previously nonflickering item was probed on only 20% of trials. In the nonflicker-bias condition, a previously flickering item was probed on 20% of trials and a previously nonflickering item was probed on 80% of trials. In the biased conditions, participants were instructed that either a flickering or a nonflickering item would be probed “most of the time.”

In total, the task consisted of 480 trials and, including breaks, took approximately 60 minutes to complete. Some participants completed the three task conditions on different days, to avoid fatigue.

Data Analysis

Performance data were converted to Cowan’s K (28), a measure of the number of items encoded in short-term memory that is more linearly related to the amount of information available than the percentage of correct responses. Conversions were performed by the following procedure: responding “change” to a change trial was considered a hit; responding “no change” to a no-change trial was considered a correct rejection; responding “change” to a no-change trial was considered a false alarm; and responding “no change” to a change trial was considered a miss. Based on these values, the hit rate [hits / (hits + misses)] and false alarm rate [false alarms / (false alarms + correct rejections)] were calculated. Hit and false alarm rates are presented in Table S1 in Supplement 1. Cowan’s K was derived by subtracting the false alarm rate from the hit rate and multiplying the result by the number of items in the tested set. This was done separately for trials in which memory was tested for a flickering versus nonflickering item to obtain separate measures of the number of flickering items and the number of nonflickering items that were stored in memory. Consequently, there were three items in the set that was tested on a given trial.

The K values were analyzed by a three-factor analysis of variance (ANOVA) with group (PSZ vs. HC) as a between-subject factor and task (unbiased, flicker-bias, nonflicker-bias) and stimulus (flickering vs. nonflickering during encoding) as within-subject factors. A significant three-way interaction was followed up by two-factor ANOVA and paired t tests. An analysis of the order of the flicker-bias and nonflicker-bias conditions is described in Supplement 1. The order of testing did not differentially affect the attentional bias effect in the two conditions or groups. Thus, carryover effects cannot explain the observed pattern of effects.

To test whether the ability to select task-relevant stimuli is related to WM capacity, we correlated each individual’s attentional selection effect in both biased conditions (flicker-bias: K for flickering minus nonflickering stimuli; nonflicker-bias: K for nonflickering minus flickering stimuli) with K scores derived from an independent WM task to ensure measurement independence from the attentional bias scores. This was a 60-trial change localization task, using the method of Gold et al. (experiment 5 in [18]). Participants viewed an array of four-colored squares, arranged around a central cross, for 100 msec (Figure S1 in Supplement 1). After a 900-msec delay, the four squares reappeared. The task was to mouse-click on the one square that had changed color.

Results

Figure 2 shows K values for HC and PSZ. The three graphs detail performance when flickering and nonflickering squares were probed with equal likelihood (unbiased) or when it was more likely that a flickering square (flicker-bias) or a nonflickering square (nonflicker-bias) would be probed. In the unbiased
condition, there was a slight competitive advantage for the flickering stimuli in both subject groups. Performance differed greatly between flickering and nonflickering items in the two bias conditions. In the flicker-bias condition, both groups displayed greater accuracy for the flickering items. In the nonflicker-bias condition, HC displayed greater accuracy for nonflickering items, while no such effect was seen in PSZ.

In the ANOVA, a main effect of group \(F(1,51) = 10.2, p = .002\) reflected lower performance in PSZ across task conditions. A significant group \(\times\) stimulus \(\times\) task interaction \(F(2,102) = 6.48, p = .002\) confirmed that the effects of stimulus type depended on the task condition but differently so in the two subject groups. The three-way interaction was followed up with two-factor ANOVA (group \(\times\) stimulus) in each task condition. In the unbiased condition, there was a trend toward better performance for flickering than nonflickering stimuli \([\text{stimulus main effect: } F(1,51) = 3.89, p = .054]\) but no stimulus \(\times\) group interaction \((p > .9)\). In the flicker-bias condition, the effects of stimulus interacted with group \([F(1,51) = 4.17, p < .05]\), suggesting that although both groups performed significantly better with the flickering stimuli, this effect was more pronounced in HC \((p < .001, \text{paired } t\text{-test})\) than PSZ \((p = .002)\). The nonflicker-bias condition yielded an even stronger stimulus \(\times\) group interaction \([F(1,51) = 7.30, p < .01]\), reflecting a significant performance advantage for the nonflickering stimuli in HC \((p < .02)\) but a trend in the opposite direction in PSZ, mimicking performance in the unbiased condition.

It is interesting to note that the effect of attention in this task consisted entirely of the suppression of distractors. That is, K for the attended stimulus type in the biased conditions always was similar to the unbiased condition, but there were performance costs for the unattended stimulus type relative to the unbiased condition. Statistically, K for the nonflickering stimuli was significantly lower in the flicker-bias than unbiased condition for both HC \((p < .001)\) and PSZ \((p = .036)\), and K for the flickering stimuli was significantly lower in the nonflicker-bias than unbiased condition for HC \((p = .013)\). However, this effect was absent in PSZ \((p > .6)\). Thus, while both groups were able to bias their attention away from the nonflickering items when a flickering item was likely to be probed, HC but not PSZ were able to bias their attention away from the salient flickering stimuli when a nonflickering stimulus was likely to be probed.

To test whether an individual’s ability to selectively encode the stimuli likely to be probed was related to WM capacity, we correlated the attentional bias effect in the two biased conditions (flicker-bias: K for flickering minus nonflickering stimuli; nonflicker-bias: K for nonflickering minus flickering stimuli) with K scores derived from an independent measure of visual WM capacity (from a change localization task). Due to color blindness, this task was not performed by 1 HC and 6 PSZ, resulting in \(n = 24\) and \(n = 22\), respectively. With data collapsed across both groups, WM capacity correlated with the attentional bias effect in both the flicker-bias \((R = .30, p < .05)\) and nonflicker-bias condition \((R = .50, p < .001)\), but the correlation was significantly greater in the nonflicker-bias conditions (Fisher’s z-transformation test for difference in correlation: \(z = 2.23, p < .05\)). The same correlations were inspected in each group individually. Although none of the correlations differed significantly between groups, the difference between the flicker-bias and nonflicker-bias condition correlations appeared to be mostly fueled by the patient group. Whereas correlations in HC were nonsignificant in both the flicker-bias \((R = .23, p > .2)\) and nonflicker-bias conditions \((R = .38, p = .07)\), PSZ displayed a similarly low correlation in the flicker-bias condition \((R = .28, p > .2)\) but a more robust correlation in the nonflicker-bias condition \((R = .47, p < .05; \text{Figure } 3)\). These trends suggest that the degree to which patients are unable to select nonsalient over salient stimuli is related to their degree of WM capacity reduction. We did not replicate this pattern of correlations with working memory indexes derived from the MATRICS Consensus Cognitive Battery; performance of these tasks is heavily influenced by executive functioning and not likely to capture much variance related to capacity.

Discussion

The present study provided evidence for impaired attentional control of working memory encoding in schizophrenia under conditions of competition from salient distractor stimuli. While both PSZ and HC were able to select salient over nonsalient stimuli, HC but not PSZ were able to select nonsalient over salient stimuli. Thus, patients could efficiently implement selection when top-down control processes were bolstered by bottom-up stimuli that conferred a competitive salience advantage.

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Figure 2. The number of flickering (flicker) or nonflickering (nonflicker) task stimuli represented in working memory in the three different task conditions. In the unbiased condition, items that had been flickering or nonflickering during encoding were equally likely to be cued. In the flicker-bias condition, an item that had been flickering during encoding was probed 80% of the time. In the nonflicker-bias condition, a previously nonflickering item was probed 80% of the time. The graphs represent averages (\(\pm\) SEM) over 28 people with schizophrenia and 25 healthy control participants. Significant differences between flickering and nonflickering stimuli are marked: \(*p < .05; **p < .01; ***p < .001\), paired t tests. HC, healthy control subjects; K, working memory capacity; PSZ, people with schizophrenia.
consistent with the top-down bias. However, selection failed when attention needed to be biased away from stimuli that had a bottom-up advantage conflicting with the top-down goals. These results extend prior observations of attentional control failures at the stage of response selection and add to previous indications that deficits in the control of input selection are present in schizophrenia (15,16,29–31). In other words, PSZ have difficulty inhibiting not only prepotent response tendencies but also salient perceptual inputs.

At first sight, there was a group difference in attentional selection not only in the nonflicker-bias, but also in the flicker-bias condition, such that PSZ displayed a smaller net benefit in K for attending the flickering over nonflickering stimuli. Thus, the selection deficit may not be entirely selective for the nonflicker-bias condition. However, when put into relation with each group’s overall performance in the flicker-bias condition, the groups appeared to allocate a similar proportion of resources to the flickering items. To quantify this, we computed the percentage of the overall storage capacity that was devoted to the flickering versus the nonflickering items (K for flickering stimuli divided by the sum of K for flickering and nonflickering stimuli in the flicker-bias condition). We found that 65% of capacity was devoted to the flickering items for HC versus 61% for PSZ. Thus, the two groups devoted an approximately equal proportion of their WM capacity to the flickering items in the flicker-bias condition. In the nonflicker-bias condition, in contrast, HC devoted 57% of capacity to the nonflickering items and PSZ devoted 47% (i.e., patients devoted more capacity to the flickering items in this condition). Thus, the group difference in attentional selection was negligible in the flicker-bias condition but substantial in the nonflicker-bias condition. Unfortunately, it was not possible to perform a valid statistical analysis of these data because this sort of ratio measure leads to extreme outliers in the single-subject data when the denominator approaches zero.

The overall lower performance observed in PSZ is in agreement with substantial and consistent impairment in visuospatial WM (32,33). For example, difficulty of PSZ in maintaining precise visuospatial information in WM has been demonstrated (34). Thus, PSZ may be particularly challenged in the visuospatial domain. However, although the current task required that spatial information be encoded and maintained in WM, other physical stimulus properties (flickering) were used to define which items should be encoded. Thus, the observed selection deficits cannot be explained by deficits in the processing of visuospatial information. A possible group difference in perceptual processing that then needs to be considered is whether PSZ and HC perceived the flickering in the same way. The data show that the flickering was salient for both groups, as PSZ and HC both displayed the same trend toward better performance with flickering items in the unbiased condition.

Participants were not instructed to maintain central fixation. Due to the short (300 msec) presentation of the encoding array, there was little time to incur large group differences in the systematic exploration of the array by overt eye movements. PSZ may have displayed a greater tendency to make eye movements to the salient flickering items they had difficulty ignoring, but this would have led to a general encoding advantage for these items relative to HC, which was not observed in the unbiased or flicker-bias conditions. Nevertheless, future studies should explore such questions by employing eye-tracking techniques. Furthermore, impaired configural processing may have caused less efficient perceptual grouping of flickering and nonflickering stimuli in PSZ than HC. Such grouping results in effectively fewer individual items to be processed and facilitates selection. The literature provides mixed evidence of a grouping deficit in schizophrenia (35–41). The divergence may reflect differences in sample composition because grouping deficits appear to be particularly prominent in patients with high levels of disorganization symptoms (42). Our sample did not include any patients with disorganized schizophrenia, and given that we studied stable outpatients, we encountered low levels of these symptoms. Thus, the present sample would not be expected to display prominent grouping deficits. Moreover, the use of in-phase flickering should have made grouping of the flickering stimuli trivially easy even for individuals with grouping deficits. Furthermore, a grouping deficit would have led to impaired performance in both the flicker-bias and nonflicker-bias conditions and could not explain the observed pattern of results.

The present findings also have implications about the nature of WM deficits in schizophrenia. Individuals differ in how much task-relevant information they can temporarily store when attention is diverted away from the perceptual input (43–45). In PSZ, reduction in measures of WM capacity is a particularly robust finding (46–50). Given that WM capacity is sharply limited, the appropriate selection of task-relevant information for consolidation is critical. There is evidence that individuals displaying low capacity scores tend to be less selective during encoding and store more irrelevant information than high-capacity individuals (51). Thus, poor WM may reflect a failure of attentional selection.
of only task-relevant items rather than a reduction in storage capacity per se.

The present findings suggest that the reductions in measures of WM capacity found in schizophrenia may be due to deficits in selectively encoding task-relevant items when irrelevant items have a bottom-up salience advantage. Indeed, on an interindividual level, the ability of PSZ to bias attention away from salient task stimuli was related to WM capacity derived from an independent task. This correlation mirrors findings on antisaccade performance in healthy subjects (52) and suggests that, in schizophrenia, a reduced ability to prevent salient distractor stimuli from occupying available WM capacity may contribute to a lower capacity for task-relevant stimuli. While in the literature the finding of lower WM capacity is not restricted to task conditions with strong bottom-up competition, salient distractors may also originate from sources outside of the administered paradigm, such as the external environment or internal sources. For example, PSZ reportedly display hyperactivity during task performance in brain regions that mediate task-independent thought, and this was suggested to reflect a misdirection of attentional resources to internal events (53).

The main finding of this study was that selection deficits in schizophrenia were particularly pronounced in, and limited to, conditions in which the top-down control of attentional resource allocation was particularly challenged by salient distractors. This specific selection deficit may partially explain lower WM capacity in schizophrenia, such that irrelevant but salient stimuli occupy storage space that could otherwise be used to hold relevant information. The present results also help settle a controversy that has recently emerged in the literature regarding whether or not PSZ display deficits in attentional selection. The present findings provide evidence that the basic mechanisms involved in implementing selection are preserved, as illustrated by the clear advantage in the recall of the high-saliency stimuli when these were attended. However, attentional selection mechanisms fail when salient bottom-up competition creates high demands on top-down control over attentional resource allocation. The present findings demonstrate that this formulation of the nature of control deficits applies across different types of selection, including the selection of not only response but also perceptual input. That is, control fails in the face of strong bottom-up competition.

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